

# Predictors of treatment dropout in self-guided web-based interventions for depression: an ‘individual patient data’ meta-analysis

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**Background.** It is well known that web-based interventions can be effective treatments for depression. However, dropout rates in web-based interventions are typically high, especially in self-guided web-based interventions. Rigorous empirical evidence regarding factors influencing dropout in self-guided web-based interventions is lacking due to small study sample sizes. In this paper we examined predictors of dropout in an individual patient data meta-analysis to gain a better understanding of who may benefit from these interventions.

**Method.** A comprehensive literature search for all randomized controlled trials (RCTs) of psychotherapy for adults with depression from 2006 to January 2013 was conducted. Next, we approached authors to collect the primary data of the selected studies. Predictors of dropout, such as socio-demographic, clinical, and intervention characteristics were examined.

**Results.** Data from 2705 participants across ten RCTs of self-guided web-based interventions for depression were analysed. The multivariate analysis indicated that male gender [relative risk (RR) 1.08], lower educational level (primary education, RR 1.26) and co-morbid anxiety symptoms (RR 1.18) significantly increased the risk of dropping out, while for every additional 4 years of age, the risk of dropping out significantly decreased (RR 0.94).

**Conclusions.** Dropout can be predicted by several variables and is not randomly distributed. This knowledge may inform tailoring of online self-help interventions to prevent dropout in identified groups at risk.

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**Key words:** Adherence, depression, eHealth, self-help, treatment, treatment dropout, web-based interventions.

## Introduction

A large body of research has suggested that web-based interventions can be effective treatments for depression with comparable effect sizes to face-to-face treatments

(Spek *et al.* 2007a; Andrews *et al.* 2010; Cuijpers *et al.* 2010). Self-guided forms of web-based treatment (i.e. interventions that patients work through on their own with no guidance) do not rely on having therapists available. These interventions can be made available to a greater number of people at very low incremental cost, thus increasing access and availability. They also maintain anonymity and overcome concern about stigmatization making them more acceptable to many people.

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However, meta-analytic studies have also shown that self-guided web-based interventions (i.e. interventions that patients work through on their own without guidance) show less promising results than guided web-based interventions that are delivered with support from a coach or therapist (Gellatly *et al.* 2007; Spek *et al.* 2007a; Andersson & Cuijpers, 2009; Cuijpers *et al.* 2011; Richards & Richardson, 2012). One explanation for the difference in effectiveness between guided and unguided web-based interventions is that the human support involved in guided interventions increases treatment adherence through accountability to a coach or therapist who is seen as trustworthy, benevolent, and an expert (Mohr *et al.* 2011). Furthermore, guided web-based interventions often not only involve a supportive coach who helps participants through the program but also more often than unguided interventions include human contact before treatment (e.g. during a diagnostic interview; Johansson and Andersson, 2012) or include referral by a therapist (Marks & Cavanagh, 2009; Berger *et al.* 2011), which may add to feelings of accountability.

In line with the idea of 'supportive accountability', higher dropout rates have been found in unguided web-based interventions for depression compared to guided web-based interventions with average levels of adherence estimated at 26% in unguided interventions and 72% in guided interventions (Richards & Richardson, 2012). In addition, empirical evidence has shown that greater exposure to the intervention is related to better treatment outcomes (Donkin *et al.* 2011) suggesting that efforts to increase adherence rates in web-based interventions may lead to better outcomes. To gain a better understanding of who may benefit from unguided web-based interventions and how we can improve adherence rates, there is a strong need to identify characteristics of individuals and interventions that are related to treatment dropout, as unguided interventions are much easier to implement and less costly than guided web-based interventions.

A few studies have already investigated this issue (Christensen *et al.* 2009; Waller & Gilbody, 2009). However, studies that have been conducted so far often lack the power to find reliable effects of predictors and moderators. In the current study we bring together the data from separate studies and employ a new strategy named individual patient data (IPD) meta-analysis. IPD meta-analysis was developed to address research questions that require large sample sizes and is based on data pooled from individual randomized control trials (RCTs) (Bower *et al.* 2013). In this way it increases the power and precision to detect predictors and moderators. This study aimed to identify socio-demographic, clinical, and intervention

characteristics that predict dropout rates in self-guided web-based interventions for depression. In the context of the present paper, the term adherence is defined as the percentage of treatment modules that were completed. Dropout rate was defined as a completion rate of <75% of the intervention modules, as we considered that in most interventions the core treatment elements are administered in this part of the treatment.

## Method

### *Search strategy for identification and selection of studies*

We used an existing database of randomized trials of psychological treatments for depression. The database has already been used by several published meta-analyses (<http://www.evidencebasedpsychotherapies.org>) and its detailed description can be found elsewhere (Cuijpers *et al.* 2008). This database has been developed and is periodically updated by a comprehensive literature search of the following health-related databases: Cochrane Central Register of control trials, PubMed, PsycINFO and EMBASE from 1996 to January 2013. In these electronic searches, various key terms covering the concepts of psychotherapy and depression were used in different combinations (both MeSH terms and text words). For a detailed description of the searches the reader is referred to Cuijpers *et al.* (2008). In addition, several systematic reviews and meta-analyses in this research field have been cross-checked throughout the development of this database in order to ensure that no trials were missing. Along with the use of this database, we contacted authors and asked them to provide us with access to the datasets of trials that were not yet published.

### *Inclusion criteria for studies*

We included (a) RCTs, (b) comparing a psychological intervention, (c) delivered through the web, (d) without any form of personal guidance, (e) with a control or comparison group, (f) aimed at adults with depression (based on a clinical interview or on elevated depressive symptoms ratings on self-report measures).

### *Quality assessment*

The validity of the studies included in the present IPD meta-analysis was examined by two independent reviewers (E.K. and D.T.) according to four criteria of the Cochrane Risk of Bias assessment tool (Higgins & Green, 2011; Higgins *et al.* 2011). We tested if the allocation concealment was adequately generated (sequence generation), the allocation was sufficiently

concealed (allocation concealment), the knowledge of the allocated intervention was adequately prevented (blinding), and any incomplete outcome data were sufficiently addressed. However, we did not consider that incomplete outcome data could influence the results of the present IPD meta-analysis since the primary aim of this paper was to identify factors influencing treatment dropout. Finally, when the information that was provided in the papers did not provide sufficient details to assess quality, we contacted the primary authors to ask what procedure was actually followed and subsequently we ran sensitivity analysis based on what the papers reported. Disagreement between the reviewers was resolved through discussion, and if needed a third reviewer was consulted (P.C.).

### Data extraction and preparation

Two authors independently extracted data included in the present meta-analysis (E.K. and D.T.). We first contacted authors of RCTs that satisfied the inclusion criteria and we asked them whether they would permit us access to their primary datasets. We identified the variables, which were common to all or most of the included datasets. These were the following: randomized group (therapy or control), baseline and follow-up depression scores, age, gender, educational level, employment status, relationship status (being in or not in a relationship), number of modules completed and presence of anxiety symptoms at baseline (yes/no; based on a clinical interview or on elevated anxiety symptoms ratings on self-report measures). Finally, we combined the individual datasets into one large pooled dataset.

### Statistical analysis

In this paper, data were extracted only for intervention groups and not for control comparison conditions as we only looked at predictors of treatment adherence. Studies included in the present IPD meta-analysis used measures such as the Center for Epidemiologic Studies Depression Scale (CES-D; Radloff, 1977), the Beck Depression Inventory – I (BDI-I; Beck *et al.* 1961) or the Beck Depression Inventory – II (BDI-II; Beck *et al.* 1996) to monitor change in depression. These depression measures were standardized (transformed into *z* scores) across the pool of the studies. We also conducted sensitivity analyses to assess the impact of baseline severity on dropout from treatment for CES-D and BDI separately. We analysed the effects of predictors on dropout from treatment using a design-based analysis of the data to account for the clustering of participants within studies. Individual patient data were analysed by a Poisson regression model for patients nested within studies to obtain

relative risks (RR) of treatment dropout on the selected factors, adjusted for the other predictors in the Poisson model and taking into account the clustered data structure by obtaining robust (Hubert–White) standard errors based on the first-order Taylor-series linearization method as implemented in Stata version SE 12.1 (StataCorp., 2011). This methodological approach is computationally efficient in synthesizing and estimating the effect of predictors (Zou, 2004). We conducted the analysis in three steps. First we conducted a series of bivariate analyses to assess the RR of each factor at a time (the so-called ‘bivariate model’). Then we repeated the analyses with all factors simultaneously entered in the Poisson model (the so-called ‘complete model’). Last, we simplified the complete model by only retaining those factors in the model that were statistically significant by eliminating factors that were not significant (the ‘parsimonious model’). Finally, we performed sensitivity analysis to assess the impact of the included studies’ quality on dropout from treatment and we checked whether the assumption of linearity was met for the relationship between dependent and independent variables.

### Ethical standards

The authors assert that all procedures contributing to this work comply with the ethical standards of the relevant national and institutional committees on human experimentation and with the Helsinki Declaration of 1975, as revised in 2008.

### Results

#### Selection of included studies

A total number of 14 164 abstracts were identified through bibliographic database searching. After the removal of duplicates, 10 474 abstracts were examined. A total of 1476 full text papers were retrieved for potential inclusion. After the exclusion of 1123 studies, 353 trials were included in the database. We searched through this database and in additional sources (grey literature, researchers on this field) and we identified 13 eligible RCTs for inclusion in the current meta-analysis. We were unable to retrieve the data from three studies (Clarke *et al.* 2002, 2005, 2009) and included 10 RCTs in the present IPD meta-analysis (77%). Fig. 1 presents the study selection process. Overall, the three studies that we did not include ( $n = 302$ ) were very similar to the 10 included studies, except for the method of recruitment (all participants in these three studies were recruited through a health mental organization in the Northwest United States, while none of the other studies recruited patients this way). The main outcome

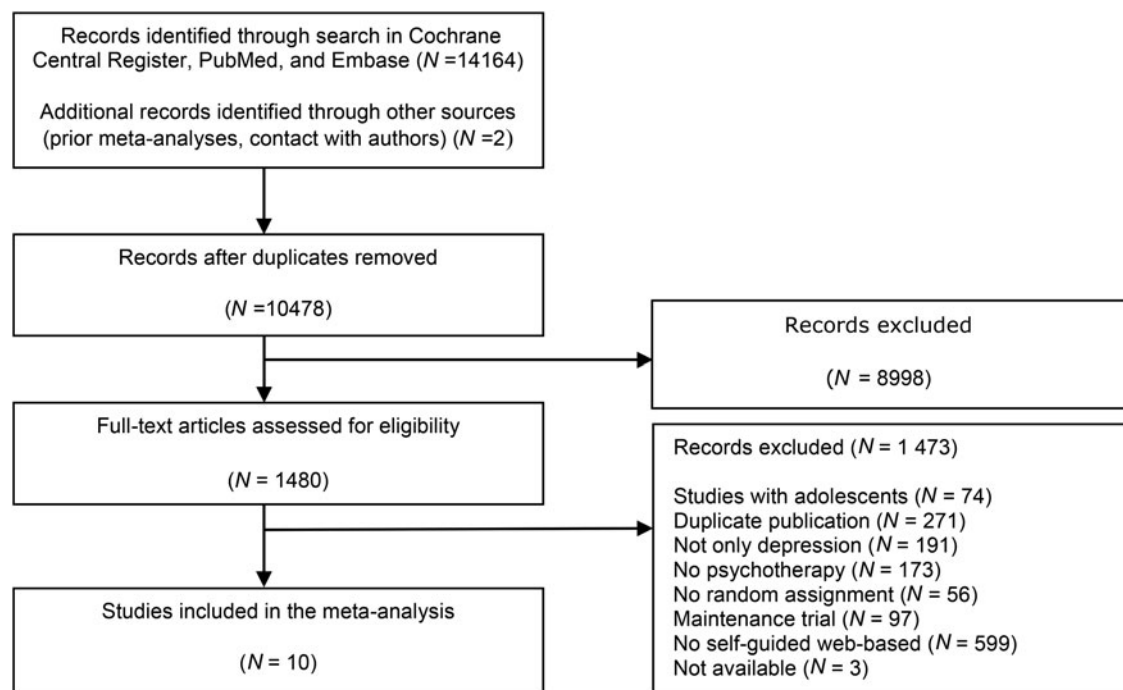


Fig. 1. Study selection process.

measures in these studies were the CES-D and the Patient Health Questionnaire-8 items (PHQ-8; Clarke et al. 2002, 2005, 2009).

### Study characteristics

In the present IPD meta-analysis, ten studies with a total of 2705 participants were included. All the examined studies recruited their participants from the community, and they were conducted across six different countries: Australia ( $n=2$ ), Germany ( $n=2$ ), Spain ( $n=1$ ), Sweden ( $n=1$ ), Switzerland ( $n=1$ ), and The Netherlands ( $n=3$ ). The majority of the included studies used self-report outcome measures for depression on which the participants needed to score above a predetermined cut-off point in order to be included in the trial. Seven out of the ten included studies used either the BDI-I or BDI-II as a primary outcome measure while the remaining trials used the CES-D.

All included unguided web-based interventions were based on three different theoretical models of psychotherapy. The majority of the included studies used interventions based on cognitive behaviour therapy (CBT) principles ( $n=8$ ). The remaining studies used either problem solving therapy (PST) ( $n=1$ ) or they compared web-based CBT with interpersonal psychotherapy (IPT) ( $n=1$ ). Table 1 shows selected characteristics of the included studies.

Most of the participants were female ( $n=1945/2705$ , 72%) and most were educated to university level

( $n=1933/2705$ , 71%). The modal age group into which participants fell was 25–34 years ( $n=741/2705$ , 27%). The average score on the CES-D at baseline assessment was 35.5 (s.d.=11.5), while the average score on the BDI was 28.4 (s.d.=13.5) indicating a high degree of severity. The average score on the CES-D and the BDI reduced at the post-treatment assessment to 24.2 (s.d.=13.2,  $n=650$ ) and 20.7 (s.d.=14.8,  $n=495$ ), respectively. The majority of the sample reported symptoms of co-morbid anxiety ( $n=1689/2705$ , 71.6%) (Table 2).

Across the 10 included RCTs, 1090 participants (40%) dropped out before the completion of 25% treatment modules, 1604 (59%) dropped out before completing half of the treatment modules. Further, levels of dropout increased to 70% (1880/2705) when we looked at the number of participants that completed 75% of treatment modules. Finally, only a small percentage (17%, 452/2705) completed all treatment modules.

### Quality assessment

All the included studies had acceptable methodological quality. The sequence was adequately generated, and the allocation was adequately concealed. Moreover, all trials used self-report outcome measures, which were administered via the Internet. Therefore, blinding of assessors was considered as adequately addressed across the ten studies of this IPD

**Table 1.** Characteristics of included studies

Study	Inclusion criteria	N	Outcome measure	Average no. of modules completed/total no. of modules	Intervention	Quality assessment <sup>a</sup>	Country
Berger <i>et al.</i> (2011)	BDI-II > 13, MDD (Mini-DIPS)	25	BDI-II	7/10	CBT	++++	Switzerland
Botella <i>et al.</i> (under submission)	Age 18–65 years, BDI-II not > 28	36	BDI-II	7/8	CBT	++++	Spain
de Graaf <i>et al.</i> (2009, 2011)	Age 8–65 years, BDI score ≥ 16	100	BDI-II	3/8	CBT	++++	The Netherlands
Donker <i>et al.</i> (2013)	CES-D ≥ 27	1864	CES-D	1/4 (CBT), 2/4 (IPT)	CBT, IPT	++++	Australia
Farrer <i>et al.</i> (2011)	K10 > 20	38	CES-D	2/5	CBT	++++	Australia
Kleiboer <i>et al.</i> (under submission)	35 > CES-D > 16; 15 > HADS > 8	107	CES-D	2/5	PST	++++	The Netherlands
Meyer <i>et al.</i> (2009)	Depression (BDI)	320	BDI	4/10	CBT	++++	Germany
Moritz <i>et al.</i> (2012)	Age 18–65 years, depression (BDI)	105	BDI	6/10	CBT	++++	Germany
Spek <i>et al.</i> (2008, 2007b)	Age 50–57 years, EDS > 12	102	BDI-II	5/10	CBT	++++	The Netherlands
Vernmark <i>et al.</i> (2010)	MDD (SCID-I-CV)	24	BDI	7/8	CBT	++++	Sweden

BDI, Beck Depression Inventory; CBT, cognitive and behavioural therapy; CES-D, Centre of Epidemiological Studies for Depression Scale; EDS, The Edinburgh Depression Scale; HADS, Hospital Anxiety and Depression Scale; IPT, interpersonal psychotherapy; K10, Kessler Psychological Distress Scale; MDD, major depressive disorder; Mini DIPS, Mini Diagnostic Interview for Psychiatric Disorders; *n*, number; SCID-I-CV, Structural Clinical Interview for DSM-IV Axis I disorders; PST, problem solving therapy.

<sup>a</sup> A positive or a negative sign is given in this column for the following quality criteria respectively: allocation sequence, allocation concealment, blinding of assessors, and incomplete outcome data (whether or not the study used intention-to-treat analysis).

**Table 2.** Demographic and clinical characteristics of the included sample

Characteristics	
Age, 25–34 years, <i>n</i> (%)	741/2705 (27)
Gender, females, <i>n</i> (%)	1945/2705 (72)
CES-D at baseline, mean (s.d.)	35.5 (11.5)
BDI at baseline, mean (s.d.)	28.4 (13.5)
CES-D at post-treatment, mean (s.d.)	24.2 (13.2)
BDI at post-treatment, mean (s.d.)	20.7 (14.8)
Co-morbid anxiety, <i>n</i> (%)	1689/2705 (71.6)

BDI, Beck Depression Inventory; CES-D, Centre for Epidemiological Studies for Depression Scale; *n*, number of patients; s.d., standard deviation.

meta-analysis. However, the participants were not blinded to the interventions, and this may have caused bias. Finally, all included RCTs used intention-to-treat analyses including all the randomized participants in their post-treatment analyses, which indicates that incomplete outcome data were adequately addressed (see Table 1).

### Predictors of dropout in self-guided web-based treatment for depression

The results of the bivariate analyses indicated that male gender (RR 1.05, 95% CI 1.01–1.11), participants with a low educational background (primary education: RR 1.23, 95% CI 1.13–1.33), the presence of co-morbid anxiety symptoms (RR 1.18, 95% CI 1.01–1.38) and CBT-based interventions (RR 1.19, 95% CI 1.03–1.39) were related to a higher risk of dropping out. Finally, the chance of dropping out significantly decreased for every 4 years of age increase (RR 0.98, 95% CI 0.97–99). The remaining variables/potential predictors (baseline severity of depression, relationship status, number of intervention modules and employment status) failed to achieve a statistically significant level of  $p < 0.05$  in the bivariate analysis (see Table 3).

Additionally, under the parsimonious model, male gender (RR 1.08, 95% CI 1.03–1.13), lower educational level (primary education: RR 1.26, 95% CI 1.14–1.39), older age (RR 0.94, 95% CI 0.87–1.02) and co-morbid anxiety (RR 1.18, 95% CI 1.01–1.38) remained statistically significant predictors of dropout from treatment. However, in our sample, CBT/not-CBT intervention status was confounded with number of modules. Thus, the effects of intervention type could not be disentangled from the number of modules and we excluded these predictors from the parsimonious model. Finally, depression severity, employment status

**Table 3.** Predictors to self-guided web-based psychotherapy for depression

Predictors	No. of studies	<i>N</i> <sub>obs</sub>	Bivariate analyses		Multivariate model ( <i>N</i> <sub>obs</sub> = 411)			Parsimonious model ( <i>N</i> <sub>obs</sub> = 2355)			
			RR	95% CI	<i>p</i>	RR	95% CI	<i>p</i>	RR	95% CI	<i>p</i>
Gender (male)	10	2702	1.05	1.01–1.11	0.04	1.04	0.82–1.31	0.72	1.08	1.03–1.13	0.002
Age (continuous per 4-year increase)	10	2700	0.98	0.97–0.99	0.001	1.01	0.96–1.07	0.05	0.98	0.97–0.99	0.004
Education											
Low <i>v.</i> high	10	2695	1.23	1.13–1.33	0.000	1.64	1.11–2.43	0.01	1.26	1.14–1.39	0.000
Middle <i>v.</i> high	10	2695	0.98	0.88–1.10	0.85	1.19	0.97–1.48	0.09	0.94	0.87–1.02	0.08
Being in a relationship (yes/no)	8	751	1.09	0.95–1.25	0.20	1.08	0.82–1.43	0.05	–	–	–
Employed (yes/no)	8	748	0.97	0.79–1.20	0.82	1.08	0.97–1.48	0.05	–	–	–
Baseline severity of depression	10	2690	1.05	0.99–1.11	0.07	1.24	0.98–1.57	0.06	–	–	–
Co-morbid anxiety	9	2358	0.09	1.01–1.38	0.02	1.36	1.01–1.82	0.04	1.18	1.01–1.38	0.03
Type of intervention (CBT <i>v.</i> others)	10	2705	1.19	1.03–1.39	0.001	0.7	0.53–1.02	0.07	–	–	–
No. of modules ( <i>n</i> < 5)	10	2705	0.88	0.72–1.08	0.24	–	–	–	–	–	–

CBT *v.* other, Cognitive behavioural therapy compared to other types of psychotherapy; CI, confidence intervals; *N*<sub>obs</sub>, number of observations; RR, risk ratio; substandard error.

and relationship status remained non-significant after controlling for the other predictors (see Table 3).

### Sensitivity analysis

We analysed the impact of depression severity on dropout for CES-D and BDI scores separately. Individuals who scored higher on CES-D at baseline had a slightly higher risk of dropping out than those with lower scores (RR 1.004, 95% CI 1.003–1.005, observations=1987,  $p < 0.001$ ). However, the increase in risk was quite small and is therefore unlikely to have clinical relevance. Further, separate analysis for BDI scores at baseline did not produce statistically significant results ( $p > 0.05$ ). It is important to stress that the results of BDI analysis were based on a considerably smaller, although sufficiently powered, number of participants ( $n = 718$ ).

Three studies did not report all relevant information regarding allocation concealment in the published papers (although personal contact with the primary authors illustrated that the allocation was adequately concealed) and thus, we decided to run sensitivity analysis based on what the papers reported. We examined the impact of quality of the included studies on treatment dropout. Study quality did not significantly predict treatment dropout ( $p > 0.05$ ). Further, we controlled for study quality in our final parsimonious model. The predictors remained the same after adjusting for the quality of the included studies.

## Discussion

### Main findings

The present IPD meta-analysis aimed to identify predictors of treatment dropout in self-guided web-based interventions for depression. We tested the relationship between dropout and several socio-demographic, clinical and intervention characteristics. The multivariate analysis of 2705 individual patients' data revealed that being male; having attained a lower educational level (primary education); a younger age and having co-morbid anxiety symptoms significantly increased the risk of dropping out before the completion of 75% of treatment modules and thus were related to high treatment's dropout.

### Placing our findings in the wider context of the literature

The finding that gender predicted treatment dropout has not been identified by previous literature on self-guided web-based interventions. However, this result may reflect a different coping strategy between the two genders. Previous research has shown that females

generally present with a higher effort to cope with depression compared to males (Babwah *et al.* 2006). These efforts might enhance their willingness to continue and complete web-based interventions without any form of guidance. There is also evidence to support the idea that women are generally more conscientious regarding health issues compared to men (Babwah *et al.* 2006). These differences in health attitudes could partly account for the differences in treatment compliance rates between the genders.

A lower educational background has also been identified as a risk for dropping out in previous research and it has been suggested that low educational status is a barrier to adherence to web-based CBT because of greater difficulties in understanding the intervention content and procedure and limited abilities in using information technology which may result in diminished motivation to continue and complete a self-guided web-based treatment (Waller & Gilbody, 2009).

Unlike the results from this study that showed that younger age was related to low treatment adherence, previous research showed that younger individuals had higher adherence to web-based treatment (Christensen *et al.* 2009).

Co-morbid anxiety symptoms increased the risk of dropping out of the treatment 16%. It is important, however, to stress that studies included in this meta-analysis were not designed for the treatment of anxiety or to deal with co-morbid anxiety and therefore the reason for this finding is unclear. Further research is needed to clarify this.

None of the remaining variables significantly predicted treatment attrition and results derived by the present IPD meta-analysis were not influenced by quality of the included studies. The lack of a significant effect on adherence of relationship status is consistent with results reported by Christensen *et al.* (2009). Further, Christensen *et al.* (2009) concluded that dropout increases with the severity of baseline depression. The findings from the present study suggest that the severity of depression does not significantly predict dropout from treatment. However, when we examined the impact of baseline severity separately for CES-D and BDI we found a significant but small higher risk for dropping out of treatment for patients who scored higher on CES-D at baseline, a result which is consistent with the conclusions of Christensen *et al.* (2009).

### Strengths and limitations

One of the strengths of the present IPD meta-analysis was that it was based on a novel methodological approach that it is considered a gold standard for identifying predictors, moderators and mediators to treatment dropout and outcome. Combining raw

individual data from several studies into one single dataset provides adequate power and precision to detect predictors of treatment attrition. Further, the systematic literature search employed by the present IPD meta-analysis reduced the risk of introducing study selection bias into the results.

In spite of the aforementioned strengths it should be noted that the present study has several limitations. Among these limitations was the risk of availability bias. We could access ten RCTs' individual patient datasets out of 13 eligible studies. Although this is higher than in other IPD meta-analyses (Riley *et al.* 2007), the ten available RCTs might differ in several ways from the three unavailable studies. Moreover, some of the predictor variables were not reported across all the ten RCTs. This might have resulted in lower power to predict effects for some of the variables of interest, although the IPD was better powered to detect a true effect than a single trial. However, such small effects would be less relevant from a clinical or public health perspective.

Moreover, the participants of the present IPD meta-analysis differ from patients in clinical samples. For instance, all the participants were recruited through the community and were proactively seeking help for their symptoms. Thus, the present findings might not be generalized to the whole population with depression but it is representative for help seeking individuals in the community. It should also be borne in mind that four of the included studies conducted a diagnostic interview before inclusion of the participants. (Spek *et al.* 2007b; de Graaf *et al.* 2009; Vernmark *et al.* 2010; Berger *et al.* 2011). This may have enhanced treatment adherence by increasing any feeling of accountability. However, we considered that these studies should be retained in our analyses since they did not provide any guidance throughout treatment. Finally, in the available data intervention type (CBT *v.* others) is confounded with the number of sessions and thus, it is not possible to reliably attribute dropout to a particular type of intervention or to the number of treatment modules.

All these predictors should be taken into account in future development of self-guided interventions for depression. For example, different features of web-based interventions may be appealing to different individuals and it is important to find out what works best for whom. Future interventions could, for example, employ more audio-visual components such as videos or gaming and less written material for individuals with a lower education. This knowledge will help in utilizing the self-guided form of web-based interventions in the most efficient and effective way. Future studies may need to be tailored to the particular needs of individuals with co-morbid anxiety

symptoms, male gender, with a low educational background and young age. Further, future research should also examine dropout at different time points or as a function of exposure to particular types of content, as treatment dropout at different time intervals may represent different processes. Other psychological predictors such as personality styles, motivation and preferences should be included in future trials to inform tailoring. This might prevent dropout in future versions of self-guided web based interventions.

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## Declaration of Interest

None.

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